

# 情報幾何学と統計多様体の幾何学

Information geometry and geometry of statistical manifolds

松添博 ( MATSUZOE Hiroshi )

名古屋工業大学大学院工学研究科 しくみ領域

0 簡単な問題設定

I 統計モデル

II 統計的推論

III ベイズ推論

IV 統計多様体と等積構造

V 不可積分系のダイバージェンス

前半の内容は 竹内純一氏 ( 九州大システム情報科学 )  
甘利俊一氏 ( 理化学研究所 ) との共同研究

## Example (Bernoulli Trial)

$$\Omega = \{0, 1\}$$

$x = 1$  : success event       $x = 0$  : failure event

$\eta$  : success probability ( $1 - \eta$  : failure probability)

$$p(x; \eta) = \eta^x (1 - \eta)^{1-x} \quad \text{the Bernoulli distribution}$$

Suppose that  $\eta$  is unknown.

Let us infer the parameter  $\eta$  from experiments (trials).

$$\text{trials: 500, success events: 298} \implies \eta = \frac{298}{500} \left( \approx \frac{3}{5} \right)$$

the maximum likelihood estimation

$$\text{trail: 1, success event: 1} \implies \eta = 1 \left( \text{We may answer } \frac{2}{3} (?), \frac{3}{4} (?) . \right)$$

## Example (Bernoulli Trial)

$$\Omega = \{0, 1\}$$

$x = 1$  : success event       $x = 0$  : failure event

$\eta$  : success probability ( $1 - \eta$  : failure probability)

$$p(x; \eta) = \eta^x (1 - \eta)^{1-x} \quad \text{the Bernoulli distribution}$$

Suppose that  $\eta$  is unknown.

Let us infer the parameter  $\eta$  from experiments (trials).

$$\text{trials: 500, success events: 298} \implies \eta = \frac{298}{500} \left( \approx \frac{3}{5} \right)$$

the maximum likelihood estimation

$$\text{trail: 1, success event: 1} \implies \eta = 1 \left( \text{We may answer } \frac{2}{3} (?), \frac{3}{4} (?) \right)$$

Bayesian estimations

We would like to consider why the ratios  $\frac{2}{3}$  or  $\frac{3}{4}$  arise  
form the viewpoint of differential geometry.

# 1 Geometry for Statistical Models

$(\Omega, \beta, P)$  : a probability space

$\Xi$  : an open domain of  $R^n$  (a parameter space)

**Definition 1.1**

$S$  is a **statistical model** or a **parametric model** on  $\Omega$

$\iff S$  is a set of probability densities with parameter  $\xi \in \Xi$  such that

$$S = \left\{ p(x; \xi) \mid \int_{\Omega} p(x; \xi) dx = 1, \quad p(x; \xi) > 0, \quad \xi \in \Xi \subset R^n \right\},$$

where  $P(A) = \int_A p(x; \xi) dx$ , ( $A \in \beta$ ).

**Example 1.2** (Normal distributions)  $\xi = (\mu, \sigma) \in \Xi = R_+^2$   
 $\mu$ : mean ( $-\infty < \mu < \infty$ ),  $\sigma$ : standard deviation ( $0 < \sigma < \infty$ ).

$$S = \left\{ p(x; \mu, \sigma) \mid p(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp \left[ -\frac{(x - \mu)^2}{2\sigma^2} \right] \right\}$$

We assume  $S$  is a smooth manifold with local coordinate system  $\Xi$ .

$g = (g_{ij})$  is the **Fisher information matrix** of  $S$

$$\begin{aligned} \iff g_{ij}(\xi) &:= \int_{\Omega} \frac{\partial}{\partial \xi^i} \log p(x; \xi) \frac{\partial}{\partial \xi^j} \log p(x; \xi) p(x; \xi) dx \\ &= \int_{\Omega} \partial_i p_{\xi} \left( \frac{\partial_j p_{\xi}}{p_{\xi}} \right) dx = E_{\xi}[\partial_i l_{\xi} \partial_j l_{\xi}] \end{aligned}$$

For simplicity, we used following notations:

$$E_{\xi}[f] = \int_{\Omega} f(x) p(x; \xi) dx, \quad (\text{the expectation of } f(x) \text{ w.r.t. } p(x; \xi)),$$

$$l_{\xi} = l(x; \xi) = \log p(x; \xi) \quad (\text{the information of } p(x; \xi)),$$

$$\partial_i = \frac{\partial}{\partial \xi^i}.$$

We assume that  $g$  is positive definite and  $g_{ij}(\xi)$  is finite for all  $i, j, \xi$ .

$\implies$  We can define a Riemannian metric on  $S$ .  
(the **Fisher metric** on  $S$ )

$g = (g_{ij})$  is the **Fisher information matrix** of  $S$

$$\begin{aligned} \overset{\text{def}}{\iff} \quad g_{ij}(\xi) &:= \int_{\Omega} \frac{\partial}{\partial \xi^i} \log p(x; \xi) \frac{\partial}{\partial \xi^j} \log p(x; \xi) p(x; \xi) dx \\ &= \int_{\Omega} \partial_i p_{\xi} \left( \frac{\partial_j p_{\xi}}{p_{\xi}} \right) dx = E_{\xi}[\partial_i l_{\xi} \partial_j l_{\xi}] \end{aligned}$$

### Proposition 1.3

*The following conditions are equivalent.*

- (1)  $g$  is positive definite.
- (2)  $\{\partial_1 p_{\xi}, \dots, \partial_n p_{\xi}\}$  are linearly independent.
- (3)  $\{\partial_1 l_{\xi}, \dots, \partial_n l_{\xi}\}$  are linearly independent.

$$\begin{aligned} \partial_i p_{\xi} &\quad \overset{\text{def}}{\iff} \text{mixture representation,} \\ \partial_i l_{\xi} = \left( \frac{\partial_i p_{\xi}}{p_{\xi}} \right) &\quad \overset{\text{def}}{\iff} \text{exponential representation.} \end{aligned}$$

For  $\alpha \in R$ , we define the  **$\alpha$ -connection**  $\nabla^{(\alpha)}$  by the following formula:

$$\Gamma_{ij,k}^{(\alpha)}(\xi) = E_\xi \left[ \left( \partial_i \partial_j l_\xi + \frac{1-\alpha}{2} \partial_i l_\xi \partial_j l_\xi \right) (\partial_k l_\xi) \right]$$

$$g(\nabla_{\partial_i}^{(\alpha)} \partial_j, \partial_k) = \Gamma_{ij,k}^{(\alpha)}$$

We can check that  $\nabla^{(\alpha)}$  ( $\forall \alpha \in R$ ) is torsion-free and  $\nabla^{(0)}$  is the Levi-Civita connection of the Fisher metricis. On the other hand,

$\nabla^{(1)}$  : the **exponential connection**

$\nabla^{(-1)}$  : the **mixture connection**

Exponential connections and mixture connections are very useful in geometric theory of statistical inferences.

For  $\alpha \in R$ , we define the  $\alpha$ -connection  $\nabla^{(\alpha)}$  by the following formula:

$$\Gamma_{ij,k}^{(\alpha)}(\xi) = E_\xi \left[ \left( \partial_i \partial_j l_\xi + \frac{1-\alpha}{2} \partial_i l_\xi \partial_j l_\xi \right) (\partial_k l_\xi) \right]$$

$$g(\nabla_{\partial_i}^{(\alpha)} \partial_j, \partial_k) = \Gamma_{ij,k}^{(\alpha)}$$

We can check that  $\nabla^{(\alpha)}$  ( $\forall \alpha \in R$ ) is torsion-free and  $\nabla^{(0)}$  is the Levi-Civita connection of the Fisher metricis. On the other hand,

$\nabla^{(1)}$  : the exponential connection

$\nabla^{(-1)}$  : the mixture connection

$$(1) \quad Xg(Y, Z) = g(\nabla_X^{(\alpha)} Y, Z) + g(Y, \nabla_X^{(-\alpha)} Z)$$

$\nabla^{(\alpha)}$  and  $\nabla^{(-\alpha)}$  are called dual (or conjugate) with respect to  $g$

$$(2) \quad g(\nabla_X^{(\alpha)} Y, Z) = g(\nabla_X^{(0)} Y, Z) - \frac{\alpha}{2} T(X, Y, Z)$$

$$T_\xi(X, Y, Z) := E_\xi[(Xl_\xi)(Yl_\xi)(Zl_\xi)]$$

the skewness or the cubic form.

$$(3) \quad (\nabla_X^{(\alpha)} g)(Y, Z) = (\nabla_Y^{(\alpha)} g)(X, Z) = \alpha T(X, Y, Z)$$

A statistical model  $S$  is an **exponential family**

$$\overset{\text{def}}{\iff}$$

$$S = \left\{ p(x; \theta) \mid p(x; \theta) = \exp[C(x) + \theta^i F_i(x) - \psi(\theta)] \right\},$$

where  $\theta^i F_i(x) = \sum_{i=1}^n \theta^i F_i(x)$  (Einstein's convention) and

$C, F_1, \dots, F_n$  : random variables on  $\Omega$

$\psi$  : a function on the parameter space  $\Theta$

The coordinate system  $[\theta^i]$  is called the **natural parameters**.

#### Proposition 1.4

For an exponential family,

(1)  $\nabla^{(1)}$  is flat

(2)  $[\theta^i]$  is an affine coordinate, i.e.,  $\Gamma_{ij}^{(1)k} \equiv 0$

Proof:

$$\Gamma_{ij,k}^{(\alpha)}(\theta) = E_\theta \left[ \left( \partial_i \partial_j l_\theta + \frac{1-\alpha}{2} \partial_i l_\theta \partial_j l_\theta \right) (\partial_k l_\theta) \right]$$

A statistical model  $S$  is an **exponential family**

$$\overset{\text{def}}{\iff}$$

$$S = \left\{ p(x; \theta) \mid p(x; \theta) = \exp[C(x) + \theta^i F_i(x) - \psi(\theta)] \right\},$$

where  $\theta^i F_i(x) = \sum_{i=1}^n \theta^i F_i(x)$  (Einstein's convention) and

$C, F_1, \dots, F_n$  : random variables on  $\Omega$

$\psi$  : a function on the parameter space  $\Theta$

The coordinate system  $[\theta^i]$  is called the **natural parameters**.

For simplicity, assume that  $C = 0$ .

### Definition 1.5

$M$  is a **curved exponential family** of  $S$

$\overset{\text{def}}{\iff}$   $M$  is a submanifold of  $S$  such that

$$M = \{p(x; \theta(u)) \mid p(x; \theta(u)) \in S, u \in U \subset \mathbb{R}^m\}$$

## Normal distributions

$\Omega = \mathbb{R}$ ,  $n = 2$ ,  $\xi = (\mu, \sigma) \in \mathbb{R}_+^2$  (the upper half plane).

$$S = \left\{ p(x; \mu, \sigma) \mid p(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{(x - u)^2}{2\sigma^2}\right] \right\}$$

The Fisher metric is

$$(g_{ij}) = \frac{1}{\sigma^2} \begin{pmatrix} 1 & 0 \\ 0 & 2 \end{pmatrix} \quad \left( S \text{ is a space of constant negative curvature } -\frac{1}{2} \right).$$

$\nabla^{(1)}$  and  $\nabla^{(-1)}$  are flat affine connections. In addition,

$$\theta^1 = \frac{\mu}{\sigma^2}, \quad \theta^2 = -\frac{1}{2\sigma^2} \quad \psi(\theta) = -\frac{(\theta^1)^2}{4\theta^2} + \frac{1}{2} \log\left(-\frac{\pi}{\theta^2}\right)$$

$$\implies p(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{(x - u)^2}{2\sigma^2}\right] = \exp[x\theta^1 + (x)^2\theta^2 - \psi(\theta)]$$

$\{\theta^1, \theta^2\}$ : natural parameters. ( $\nabla^{(1)}$ -geodesic coordinate system)

$$\eta_1 = E[x] = \mu, \quad \eta_2 = E[x^2] = \sigma^2 + \mu^2.$$

$\{\eta_1, \eta_2\}$ : moment parameters. ( $\nabla^{(-1)}$ -geodesic coordinate system)

## Finite sample space

$$\Omega = \{x_0, x_1, \dots, x_n\}, \dim S = n$$

$$p(x_i; \eta) = \begin{cases} \frac{\eta_i}{1 - \sum_{j=1}^n \eta_j} & (1 \leq i \leq n) \\ 1 - \sum_{j=1}^n \eta_j & (i = 0) \end{cases}$$

$$\Xi = \left\{ \{\eta_1, \dots, \eta_n\} \mid \eta_i > 0 \ (\forall i), \sum_{j=1}^n \eta_j < 1 \right\}$$

(an  $n$ -dimensional simplex)

The Fisher metric is

$$(g_{ij}) = \frac{1}{\eta_0} \begin{pmatrix} 1 + \frac{\eta_0}{\eta_1} & 1 & \cdots & 1 \\ 1 & 1 + \frac{\eta_0}{\eta_2} & & \vdots \\ \vdots & & \ddots & \vdots \\ 1 & \cdots & \cdots & 1 + \frac{\eta_0}{\eta_n} \end{pmatrix},$$

$$\text{where } \eta_0 = 1 - \sum_{j=1}^n \eta_j.$$

$\left( \textcolor{blue}{S \text{ is a space of constant positive curvature } \frac{1}{4}} \right).$

## Finite sample space

$$\Omega = \{x_0, x_1, \dots, x_n\}, \dim S = n$$

$$p(x_i; \eta) = \begin{cases} \frac{\eta_i}{1 - \sum_{j=1}^n \eta_j} & (1 \leq i \leq n) \\ 1 - \sum_{j=1}^n \eta_j & (i = 0) \end{cases}$$

$$\Xi = \left\{ \{\eta_1, \dots, \eta_n\} \mid \eta_i > 0 \ (\forall i), \sum_{j=1}^n \eta_j < 1 \right\}$$

(an  $n$ -dimensional simplex)

$\{\theta^1, \dots, \theta^n\}$ : natural parameters. ( $\nabla^{(1)}$ -geodesic coordinate system)

$$\text{where } \theta^i = \log \frac{\eta_i}{1 - \sum_{j=1}^n \eta_j} = \log \frac{p(x_i)}{p(x_0)}.$$

$\{\eta_1, \dots, \eta_n\}$ : moment parameters. ( $\nabla^{(-1)}$ -geodesic coordinate system)

## Bernoulli distributions

$$\Omega = \{0, 1\}, n = 1, \xi = \eta.$$

$$C(x) = 0, \quad F(x) = x, \quad \theta = \log \frac{\eta}{1 - \eta},$$

$$\psi(\theta) = -\log(1 - \eta) = \log(1 + e^\theta)$$

Then we obtain

$$p(x; \xi) = \eta^x (1 - \eta)^{1-x} = \exp [\log \eta^x (1 - \eta)^{1-x}]$$

$$= \exp [x\theta - \psi(\theta)].$$

This implies that Bernoulli distributions are an exponential family.  
The expectation parameter is:

$$E[x] = 1 \cdot \eta + 0 \cdot (1 - \eta) = \eta$$

The Fisher metric is

$$g(\eta) = \frac{1}{\eta(1 - \eta)}$$

## 0 簡単な問題設定

I 統計モデル

II 統計的推論

III ベイズ推論

IV 統計多様体と等積構造

V 不可積分系のダイバージェンス

## 2 Statistical inference for curved exponential families

$S$  : an exponential family

$M$  : a curved exponential family embedded into  $S$

$x_1, \dots, x_N$  :  $N$  independent observations of the random variable  $x$   
distributed to  $p(x; u) \in M$

Given  $x^N = (x_1, \dots, x_N)$ , a function  $L$  on  $U$  can be defined by

$$\begin{aligned} L(u) &= p(x_1; u) \cdots p(x_N; u) \\ &= \prod_{i=1}^n p(x_i; u) \\ &= p(x^N; u) \end{aligned}$$

We call  $L$  a **likelihood function**.

We say that a statistic is the **maximum likelihood estimator** if it maximizes the likelihood function:

$$\hat{u} = \arg \max_{u \in U} L(u), \quad \left( L(\hat{u}) = \max_{u \in U} L(u) \right)$$

Suppose that  $p(x; \theta), p(x; \theta') \in S$ .

$KL$  : the **Kullback-Leibler divergence**  
(or the **relative entropy**) of  $S$

$\overset{\text{def}}{\iff}$   $KL$  is a function on  $S \times S$  such that

$$KL(p(\theta) || p(\theta')) = \int_{\Omega} \log \frac{p(\theta)}{p(\theta')} p(\theta) dx.$$

$$\bar{x} = \frac{1}{N} \sum x_i \quad (\text{the sample mean of } x^N)$$

$$\hat{\eta}_i = \frac{1}{N} \sum_{j=1}^N F_i(x_j) \quad (\text{the sample mean of the random variable } F_i.)$$

$$\phi(\theta) = E_{\theta}[\log p(\theta)] \quad (-\phi(\theta) \text{ is the entropy of } p(\theta))$$

Then the Kullback-Leibler divergence is given by

$$KL(p(\hat{\eta}) || p(u)) = \phi(\hat{\eta}) - \frac{1}{N} \log L(u).$$

The maximum likelihood estimation  $\hat{u}$  is the point in  $M$  which minimizes the divergence from  $p(\hat{\eta})$ .

---

## KL-divergence (statistically)

---

the Kullback-Leibler divergence

$$\begin{aligned} KL(p(\theta) || p(\theta')) &= \int_{\Omega} \log \frac{p(\theta)}{p(\theta')} p(\theta) dx \\ &= \int_{\Omega} (\log p(\theta) - \log p(\theta')) p(\theta) dx \end{aligned}$$

The Kullback-Leibler divergence measures the difference of the mean of informations from  $\log p(\theta)$  to  $\log p(\theta')$ .

---

## KL-divergence (geometrically)

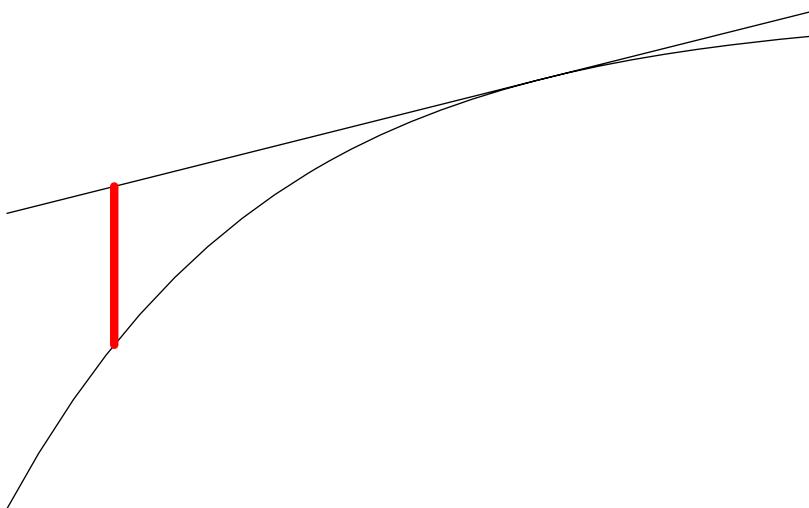
---

Suppose that  $M$  is an exponential family.

$$\begin{array}{ll} \phi(\theta) = E_{\theta}[\log p(\theta)] & (-\phi(\theta)) \text{ is the entropy of } p(\theta) \\ l_{\theta} & : \text{the tangent hyperplane of } \phi \text{ at } \theta \end{array}$$

$$KL(p(\theta) || p(\theta')) = l_{\theta}(\theta') - \phi(\theta')$$

The Kullback-Leibler divergence measures the difference of the height between  $l_{\theta}(\theta')$  and  $\phi(\theta')$ .



The KL-divergence  $KL(p||q)$  is equivalent to the difference between  $\phi(\theta')$  and  $l_\theta(\theta')$ .

This implies the KL-divergence is contained in the class of Bragman divergences (or canonical divergences).

### — KL-divergence (geometrically) —

Suppose that  $M$  is an exponential family.

$$\phi(\theta) = E_\theta[\log p(\theta)] \quad (-\phi(\theta) \text{ is the entropy of } p(\theta))$$

$l_\theta$  : the tangent hyperplane of  $\phi$  at  $\theta$

$$KL(p(\theta)||p(\theta')) = l_\theta(\theta') - \phi(\theta')$$

The Kullback-Leibler divergence measures the difference of the height between  $l_\theta(\theta')$  and  $\phi(\theta')$ .

## 0 簡単な問題設定

I 統計モデル

II 統計的推論

III ベイズ推論

IV 統計多様体と等積構造

V 不可積分系のダイバージェンス

### 3 Bayesian inference of curved exponential families

$S$  : an exponential family

$M$  : a curved exponential family embedded into  $S$

$p(x; \theta(u))$  : the model distribution which generates data

$\rho(u)du$  : a prior distribution

e.g.  $\tilde{\rho}^{(0)}$  : the Jeffreys prior of  $M$ .

$$\overset{\text{def}}{\iff} \tilde{\rho}^{(0)} = \frac{(\det |g_{ab}|)^{1/2}}{\int_U (\det |g_{ab}|)^{1/2} du} du \quad g : \text{the Fisher metric of } M.$$

We define the posterior distribution by

$$\rho'(u|x) = \frac{p(x; u)\rho(u)}{\int_U p(x; u)\rho(u)du}.$$

$x^N$  :  $N$  observations obtained from  $p(x; \theta(u))$ .

We define the Bayesian mixture distribution by

$$f_\rho[x^N](x) = \int_U p(x; u)\rho'(u|x^N)du$$

Let us consider the projection from  $f_\rho[x^N](x)$  to  $M$  with respect to the Kullback-Leibler divergence:

$$u \left( \tilde{f}_\rho[x^N] \right) = \arg \min_{u \in U} KL \left( f_\rho[x^N] || p(x^N; u) \right).$$

$u \left( \tilde{f}_\rho[x^N] \right)$  : the **projected Bayesian estimation**.

## Example (Bernoulli Trial)

$$\Omega = \{0, 1\}$$

$$p(x; \eta) = \eta^x (1 - \eta)^{1-x}$$

$\eta$  : an expectation parameter

$$\theta = \log \frac{\eta}{1 - \eta} \quad \text{a natural parameter}$$

$$g(\eta) = \frac{1}{\eta(1 - \eta)} \quad \text{the Fisher information with respect to } \eta$$

priors	$d\theta$	Jeffreys	$d\eta$
density $\rho(\eta)$ w.r.t. $d\eta$	$\frac{d\theta}{d\eta} = \frac{1}{\eta(1 - \eta)}$	$\frac{1}{\sqrt{\eta(1 - \eta)}}$	1

where  $d\theta$  and  $d\eta$  are uniform priors with respect to  $\theta$  and  $\eta$ , respectively.

## $\alpha$ -parallel priors

Recall the Bayes formula:

$$\rho'(u|x) = \frac{p(x; u)\rho(u)}{\int_U p(x; u)\rho(u)du}$$

The integral is carried out on the parameter space

$\implies$  A prior distribution can be regarded as a **volume element** on  $M$ .

$M$  : a statistical model

$g$  : the Fisher metric on  $M$

$\nabla^{(0)}$  : the Levi-Civita connection with respect to  $g$

$\tilde{\omega}^0$  : the Jeffreys prior distribution

Proposition 3.1  $\nabla^{(0)}\tilde{\omega}^0 = 0$

Definition 3.2

$\tilde{\omega}^{(\alpha)}$  is an  **$\alpha$ -(parallel) prior**  $\iff \nabla^{(\alpha)}\tilde{\omega}^{(\alpha)} = 0$

For an exponential family

$d\theta \leftrightarrow$  1-parallel prior       $d\eta \leftrightarrow -1$ -parallel prior

## Example (Bernoulli Trial)

$$\Omega = \{0, 1\}, \quad p(x; \eta) = \eta^x (1 - \eta)^{1-x}$$

$\eta$  : an expectation parameter

$\theta = \log \frac{\eta}{1 - \eta}$  a natural parameter

$g(\eta) = \frac{1}{n(1 - \eta)}$  the Fisher information with respect to  $\eta$

priors	$d\theta$	Jeffreys	$d\eta$
density $\rho(\eta)$ w.r.t. $d\eta$	$\frac{d\theta}{d\eta} = \frac{1}{\eta(1 - \eta)}$	$\frac{1}{\sqrt{\eta(1 - \eta)}}$	1

Experiment  $N = 1$ , success event  $k = 1$

	$d\theta$	Jeffreys	$d\eta$
the projected Bayes estimator	1	$\frac{3}{4}$	$\frac{2}{3}$

General case

	$\frac{k}{N}$	$\frac{k + \frac{1}{2}}{N + 1}$	$\frac{k + 1}{N + 2}$
the projected Bayes estimator			

## 0 簡単な問題設定

I 統計モデル

II 統計的推論

III ベイズ推論

IV 統計多様体と等積構造

V 不可積分系のダイバージェンス

## 4 Statistical manifolds and equiaffine structures

$(M, g)$  : a Riemannian manifold

$\nabla$  : a torsion-free affine connection on  $M$

$$\text{i.e. } T^\nabla(X, Y) := \nabla_X Y - \nabla_Y X - [X, Y] \equiv 0$$

### Definition 4.1

We call the triplet  $(M, \nabla, g)$  a **statistical manifold**

$$\overset{\text{def}}{\iff} \quad \nabla g \text{ is totally symmetric.}$$

### Definition 4.2

$\nabla^*$ : **dual (or conjugate) connection** of  $\nabla$  with respect to  $g$  by

$$Xg(Y, Z) = g(\nabla_X Y, Z) + g(Y, \nabla_X^* Z).$$

$$(1) (\nabla^*)^* = \nabla$$

$$(2) \text{ Set } \nabla^{(0)} = \frac{1}{2}(\nabla + \nabla^*) \implies \nabla^{(0)}g = 0$$

(3)  $(M, \nabla, g)$  : statistical manifold  $\iff (M, \nabla^*, g)$  : statistical manifold  
 $(M, \nabla^*, g)$  : the **dual statistical manifold**

**Proposition 4.3**

$(M, g)$  : Riemannian manifold with the Levi-Civita connection  $\nabla^{(0)}$   
 $C$  : totally symmetric  $(0, 3)$ -tensor field

$$g(\nabla_X Y, Z) := g(\nabla_X^{(0)} Y, Z) - \frac{1}{2}C(X, Y, Z),$$

$$g(\nabla_X^* Y, Z) := g(\nabla_X^{(0)} Y, Z) + \frac{1}{2}C(X, Y, Z),$$

$\implies$  (1)  $\nabla$  and  $\nabla^*$  are torsion-free dual affine connections.  
(2)  $\nabla g$  and  $\nabla^* g$  are totally symmetric.

**Proposition 4.4**

If we assume two conditions from the followings, then the others hold.

- (1)  $\nabla$  is torsion-free.
- (2)  $\nabla^*$  is torsion-free.
- (3)  $C = \nabla g$  is totally symmetric.
- (4)  $\nabla^{(0)} = (\nabla + \nabla^*)/2$  is the Levi-Civita connection with respect to  $g$ .

**Proposition 4.3**

$(M, g)$  : Riemannian manifold with the Levi-Civita connection  $\nabla^{(0)}$   
 $C$  : totally symmetric  $(0, 3)$ -tensor field

$$g(\nabla_X Y, Z) := g(\nabla_X^{(0)} Y, Z) - \frac{1}{2}C(X, Y, Z),$$

$$g(\nabla_X^* Y, Z) := g(\nabla_X^{(0)} Y, Z) + \frac{1}{2}C(X, Y, Z),$$

$\implies$  (1)  $\nabla$  and  $\nabla^*$  are torsion-free dual affine connections.  
(2)  $\nabla g$  and  $\nabla^* g$  are totally symmetric.

**Definition 4.1 (Kurose)**

We call the triplet  $(M, \nabla, g)$  a **statistical manifold**

$\overset{\text{def}}{\iff}$   $\nabla g$  is totally symmetric.

**Definition 4.5 (Lauritzen)**

$(M, g)$  : a Riemannian manifold

$C$  : a totally symmetric  $(0, 3)$ -tensor field

We call the triplet  $(M, g, C)$  a **statistical manifold**.

## Parametric statistical model

$(\Omega, \beta, P)$  : a probability space

$\Xi$  : an open domain of  $R^n$  (a parameter space)

$S$  is a **statistical model** or a **parametric model** on  $\Omega$

$\iff$   $S$  is a set of probability densities with parameter  $\xi \in \Xi$  such that

$$S = \left\{ p(x; \xi) \left| \int_{\Omega} p(x; \xi) dx = 1, p(x; \xi) > 0, \xi \in \Xi \subset R^n \right. \right\},$$

where  $P(A) = \int_A p(x; \xi) dx$ , ( $A \in \beta$ ).

$g = (g_{ij})$  is the **Fisher information matrix** of  $S$

$$\iff g_{ij}(\xi) := \int_{\Omega} \frac{\partial}{\partial \xi^i} \log p(x; \xi) \frac{\partial}{\partial \xi^j} \log p(x; \xi) p(x; \xi) dx (= E_{\xi}[\partial_i l_{\xi} \partial_j l_{\xi}])$$

We assume that  $g$  is positive definite and  $g_{ij}(\xi)$  is finite for all  $i, j, \xi$ .

$\implies$  We can define a Riemannian metric on  $S$ . (the **Fisher metric**)

For  $\alpha \in R$ , the  $\alpha$ -connection  $\nabla^{(\alpha)}$  on  $S$

$$\begin{aligned} &\iff \Gamma_{ij,k}^{(\alpha)}(\xi) = E_\xi \left[ \left( \partial_i \partial_j l_\xi + \frac{1-\alpha}{2} \partial_i l_\xi \partial_j l_\xi \right) (\partial_k l_\xi) \right] \\ &g(\nabla_{\partial_i}^{(\alpha)} \partial_j, \partial_k) = \Gamma_{ij,k}^{(\alpha)} \end{aligned}$$

We can check that  $\nabla^{(\alpha)}$  ( $\forall \alpha \in R$ ) is torsion-free and  $\nabla^{(0)}$  is the Levi-Civita connection of the Fisher metrics. On the other hand,

- $\nabla^{(1)}$  : the exponential connection
- $\nabla^{(-1)}$  : the mixture connection

$$(1) \quad Xg(Y, Z) = g(\nabla_X^{(\alpha)} Y, Z) + g(Y, \nabla_X^{(-\alpha)} Z)$$

$\nabla^{(\alpha)}$  and  $\nabla^{(-\alpha)}$  are called dual (or conjugate) with respect to  $g$

$$(2) \quad g(\nabla_X^{(\alpha)} Y, Z) = g(\nabla_X^{(0)} Y, Z) - \frac{\alpha}{2} T(X, Y, Z)$$

$T_\xi(X, Y, Z) := E_\xi[(Xl_\xi)(Yl_\xi)(Zl_\xi)]$  : skewness, or cubic form.

$$(3) \quad (\nabla_X^{(\alpha)} g)(Y, Z) = (\nabla_Y^{(\alpha)} g)(X, Z) = \alpha T(X, Y, Z)$$

$(M, g)$  : a Riemannian manifold

$\nabla^{(0)}$  : the Levi-Civita connection with respect to  $g$

$C$  : a totally symmetric  $(0, 3)$ -tensor field on  $M$

For fixed  $\alpha \in R$ , an  **$\alpha$ -connection** is defined by

$$g(\nabla_X^{(\alpha)} Y, Z) := g(\nabla_X^{(0)} Y, Z) - \frac{\alpha}{2} C(X, Y, Z)$$

(1)  $\nabla^{(\alpha)}, \nabla^{(-\alpha)}$  are **mutually dual** torsion-free affine connections

$$Xg(Y, Z) = g(\nabla_X^{(\alpha)} Y, Z) + g(Y, \nabla_X^{(-\alpha)} Z).$$

(2)  $\alpha \in R \implies (\nabla_X^{(\alpha)} g)(Y, Z) = \alpha C(X, Y, Z)$

(3)  $(M, \nabla^{(\alpha)}, g)$  is a statistical manifold.

#### Definition 4.6

$(M, \nabla, g)$  : a statistical manifold,

$T$  : the **Tchebychev form**, and  $\#T$  : the **Tchebychev vector field**

$$\begin{aligned} &\iff T(X) := \text{trace}_g\{(Y, Z) \mapsto C(X, Y, Z)\}, \\ &g(\#T, X) := T(X) \end{aligned}$$

$M$  : an  $n$ -dimensional manifold

$\nabla$  : a torsion-free affine connection on  $M$

$\omega$  : a volume element of  $M$ ,

Definition 4.7  $\{\nabla, \omega\}$  is (locally) equiaffine structure on  $M$ .

$$\iff^{\text{def}} \nabla\omega = 0$$

$\nabla$  is called a (locally) equiaffine connection ,

$\omega$  is called a parallel volume element.

Proposition 4.8

$(M, \nabla, g)$  : a statistical manifold

$T$  : the Tchebychev form

Then  $\nabla$  is an equiaffine connection  $\iff dT = 0$

$\nabla^{(\alpha)}$  is equiaffine

$\implies dT = 0 \implies$  there exists a function  $\phi$  on  $M$  such that  $T = d\phi$ .

Hence  $g(\#T, X) = X\phi$

The Tchebychev vector field is a gradient vector field of some function  $\phi$  on  $M$ .

### Proposition 4.9

$(M, g, C)$  : a statistical manifold

$\nabla^{(\alpha)}, \nabla^{(-\alpha)}$  : affine connectoions determined by  $g, C$

$T = d\phi$  : the Tchebychev form on  $(M, g, C)$

Then

$\{\nabla^{(\alpha)}, \omega\}$  is an equiaffine structure

$\iff \{\nabla^{(-\alpha)}, e^{-\alpha\phi}\omega\}$  is an equiaffine structure.

### Theorem 4.10

$\hat{u}$  : the maximum likelihood estimator (MLE)

$\hat{g}$  : the Fisher metric with respect to MLE

$\hat{C}$  : the skewness tensor with respect to MLE

$u(\tilde{f}^{(\alpha)}[x^N])$  : the projected Bayesian estimator with  $\alpha$ -paralell prop

$$\begin{aligned} \implies u^c(\tilde{f}^{(\alpha)}[x^N]) &= \hat{u}^c + \frac{1-\alpha}{2N} \hat{C}_{abd} \hat{g}^{ab} \hat{g}^{cd} + o\left(\frac{1}{N}\right) \\ &= \hat{u}^c + \frac{1-\alpha}{2N} \# \hat{T}^c + o\left(\frac{1}{N}\right) \end{aligned}$$

## 0 簡単な問題設定

I 統計モデル

II 統計的推論

III ベイズ推論

IV 統計多様体と等積構造

V 不可積分系のダイバージェンス

## 5 Divergences for non-integrable systems

### 5.1 Dually flat spaces

$(M, g)$  : a Riemannian manifold

$\nabla$  : a torsion-free affine connection on  $M$

We call the triplet  $(M, \nabla, g)$  a **statistical manifold**  
 $\overset{\text{def}}{\iff}$   $\nabla g$  is totally symmetric.

$\nabla^*$ : **dual (or conjugate) connection** of  $\nabla$  with respect to  $g$  by

$$Xg(Y, Z) = g(\nabla_X Y, Z) + g(Y, \nabla_X^* Z).$$

$\nabla$  is flat  $\iff$   $\nabla^*$  is flat.

$(M, g, \nabla, \nabla^*)$  : **dually flat space**  $\overset{\text{def}}{\iff}$   $\nabla, \nabla^*$  are flat affine connections.

An affine connection  $\nabla$  is flat

$\implies$  there exists a local coordinate system on  $M$  such that

$$\Gamma_{ij}^{\nabla k} \equiv 0.$$

We call such a coordinate system an affine coordinate system.

**Proposition 5.1**

$(M, g, \nabla, \nabla^*)$  : dually flat space

$\{\theta^i\}$  :  $\nabla$ -affine coordinate system

$\implies$  there exists an  $\nabla^*$ -affine coordinate system  $\{\eta_i\}$  such that

$$g\left(\frac{\partial}{\partial \theta^i}, \frac{\partial}{\partial \eta_j}\right) = \delta_i^j.$$

$\{\eta_i\}$  : the dual coordinate system with respect to  $\{\theta^i\}$ .

**Proposition 5.2**

$(M, g, \nabla, \nabla^*)$  : a dually flat space

$\{\theta^i\}$  : a  $\nabla$ -affine coordinate system

$\{\eta_i\}$  : the dual coordinate system of  $\{\theta^i\}$

$\implies$  there exists functions  $\psi, \phi$  on  $M$  such that

$$\frac{\partial \psi}{\partial \theta^i} = \eta_i, \quad \frac{\partial \phi}{\partial \eta_i} = \theta^i, \quad \psi(p) + \phi(p) - \sum_{i=1}^m \theta^i(p) \eta_i(p) = 0. \quad (1)$$

In addition, the following formulas hold

$$g_{ij} = \frac{\partial^2 \psi}{\partial \theta^i \partial \theta^j}, \quad g^{ij} = \frac{\partial^2 \phi}{\partial \eta_i \partial \eta_j}, \quad (2)$$

where

$(g_{ij})$  : the component matrix of a Riemannian metric  $g$ ,

$(g^{ij})$  : the inverse matrix of  $(g_{ij})$

$\psi$  is the  **$\theta$ -potential** function, and  $\phi$  is the  **$\eta$ -potential** function.

We say that the relation (1) the **Legendre transformation**.

### Definition 5.3

$\rho : M \times M \rightarrow R$  : (canonical) divergence on  $(M, g, \nabla, \nabla^*)$

$$\overset{\text{def}}{\iff} \quad \rho(p||q) := \psi(p) + \phi(q) - \sum_{i=1}^n \theta^i(p)\eta_i(q), \quad (p, q \in M).$$

### Proposition 5.4

*The definition of  $\rho$  is independent of choice of affine coordinate system on  $M$ .*

### Example 5.5 (Euclidean space)

$R^m$  : Euclidean space

$\langle , \rangle$  : the standard inner product

$D$  : the standard flat affine connection.

$\implies (M, \langle , \rangle, D, D)$  is a dually flat space,

$$\rho(p||q) = \frac{1}{2}d(p, q)^2.$$

---

## KL-divergence

---

the Kullback-Leibler divergence

$$\begin{aligned}
 KL(p(\theta) || p(\theta')) &= \int_{\Omega} \log \frac{p(\theta)}{p(\theta')} p(\theta) dx \\
 &= \int_{\Omega} (\log p(\theta) - \log p(\theta')) p(\theta) dx
 \end{aligned}$$

Suppose that  $S$  is an exponential family and  $p(\theta), p(\theta') \in S$ .

$$\begin{aligned}
 KL(p(\theta) || p(\theta')) &= \int_{\Omega} \left( \sum_{i=1}^n \theta^i F_i(x) - \psi(\theta) - \sum_{i=1}^n \theta'^i F_i(x) + \psi(\theta') \right) p(\theta) dx \\
 &= \psi(\theta') - \psi(\theta) + \sum_{i=1}^n \theta^i \eta_i - \sum_{i=1}^n \theta'^i \eta_i \\
 &= \psi(\theta') + \phi(\theta) - \sum_{i=1}^n \theta'^i \eta_i \\
 &= \rho(p(\theta') || p(\theta))
 \end{aligned}$$

$(M, \nabla, h)$  : a simply connected flat statistical manifold.

( $\Rightarrow (M, h, \nabla, \nabla^*)$  is a dually flat space.)

$\Rightarrow \exists \psi : \text{a function on } M \text{ (potential function) such that } \frac{\partial^2 \psi}{\partial \theta^i \partial \theta^j} = g_{ij}$

$\Rightarrow f : M \rightarrow R^{n+1} : \text{an immersion } (\{f, \xi\} \text{ a graph immersion})$

$$f : \begin{pmatrix} \theta^1 \\ \vdots \\ \theta^n \end{pmatrix} \mapsto \begin{pmatrix} \theta^1 \\ \vdots \\ \theta^n \\ \psi(\theta) \end{pmatrix}, \quad \xi = \begin{pmatrix} 0 \\ \vdots \\ 0 \\ 1 \end{pmatrix}$$

$v : M \rightarrow R^{n+1} : \text{the conormal map of } \{f, \xi\}$ ,

$$v = (-\eta_1, \dots, -\eta_n, 1) \quad \eta_i = \frac{\partial \psi}{\partial \theta^i}$$

From  $\phi(q) = \sum \eta_i(q) \theta^i(q) - \psi(q)$ , we define the **geometric divergence** by

$$\begin{aligned} \rho(p, q) &= \langle v(q), f(p) - f(q) \rangle \\ &= - \sum \eta_i(q) \theta^i(p) + \psi(p) + \sum \eta_i(q) \theta^i(q) - \psi(q) \\ &= \psi(p) + \phi(q) - \sum \eta_i(q) \theta^i(p) = \rho^C(p, q) \end{aligned}$$

The geometric divergence coincides with the canonical divergence.

$(M, \nabla, h)$  : a simply connected flat statistical manifold.

( $\Rightarrow (M, h, \nabla, \nabla^*)$  is a dually flat space.)

$\Rightarrow \exists \psi : \text{a function on } M \text{ (potential function) such that } \frac{\partial^2 \psi}{\partial \theta^i \partial \theta^j} = g_{ij}$

$\Rightarrow f : M \rightarrow R^{n+1} : \text{an immersion } (\{f, \xi\} \text{ a graph immersion})$

$$f : \begin{pmatrix} \theta^1 \\ \vdots \\ \theta^n \end{pmatrix} \mapsto \begin{pmatrix} \theta^1 \\ \vdots \\ \theta^n \\ \psi(\theta) \end{pmatrix}, \quad \xi = \begin{pmatrix} 0 \\ \vdots \\ 0 \\ 1 \end{pmatrix}$$

$v : M \rightarrow R^{n+1}$  is the **conormal map** of  $\{f, \xi\}$

$$\stackrel{\text{def}}{\iff} \begin{aligned} \langle v(p), \xi_p \rangle &= 1, \\ \langle v(p), f_* X_p \rangle &= 0 \end{aligned}$$

We define a function on  $M \times M$  by

$$\rho(p, q) = \langle v(q), f(p) - f(q) \rangle$$

$\rho$  is called the **geometric divergence** on  $M$ .

## 最尤法の復習

$S$  : an exponential family

$M$  : a curved exponential family embedded into  $S$ ,  $U \subset M$

Given  $(x_1, \dots, x_N)$ , the **likelihood function**  $L$  is defined by

$$L(u) = p(x_1; u) \cdots p(x_N; u) = \prod_{i=1}^n p(x_i; u)$$

$$\hat{u} : \text{the maximum likelihood estimator} \quad \stackrel{\text{def}}{\iff} \quad \hat{u} = \arg \max_{u \in U} L(u)$$

実際に最尤推定量を求めるためには，対数尤度方程式を解く．

$$\frac{\partial}{\partial u^1} \log L(u) = 0, \quad \dots, \quad \frac{\partial}{\partial u^m} \log L(u) = 0$$

$$\hat{\eta}_i = \frac{1}{N} \sum_{j=1}^N F_i(x_j) \quad (F_i \text{ の標本平均}), \quad \phi(\theta) = E_\theta[\log p(\theta)]$$

$$KL(p(\hat{\eta}) || p(u)) = \phi(\hat{\eta}) - \frac{1}{N} \log L(u)$$

すなわち  $\hat{u}$  : ある推定量  $\iff d(KL_{\hat{\eta}})(X_u) = 0$  ( $\forall X_u \in T_u M$ )

## 5.2 Divergences for non-integrable systems

$\omega : \Gamma(TM) \rightarrow \mathbf{R}^{n+1}$ : a  $\mathbf{R}^{n+1}$ -valued 1-form

$\xi : M \rightarrow \mathbf{R}^{n+1}$ : a  $\mathbf{R}^{n+1}$ -valued function

### Definition 5.6

$\{\omega, \xi\}$  is called an affine distribution

$\overset{\text{def}}{\iff}$  For an arbitrary point  $p \in M$ ,

- (1)  $\mathbf{R}^{n+1} = \text{Image } \omega_p \oplus \mathbf{R}\{\xi_x\}$
- (2)  $\text{Image } (d\omega)_p \subset \text{Image } \omega_p$

$$\begin{aligned} X\omega(Y) &= \omega(\nabla_X Y) + h(X, Y)\xi, \\ X\xi &= -\omega(SX) + \tau(X)\xi. \end{aligned}$$

$\omega$  : non-degenerate  $\overset{\text{def}}{\iff}$   $h$  : non-degenerate

$\{\omega, \xi\}$  : equiaffine  $\overset{\text{def}}{\iff}$   $\tau = 0$

Remark 5.7  $\text{Image } (d\omega)_p \subset \text{Image } \omega_p \iff h$ : symmetric.

## SLD Fisher metrics

$\text{Herm}(d)$  : the set of all Hermitian matrices of degree  $d$ .

$\mathcal{S}$  : a space of quantum states

$$\mathcal{S} = \{P \in \text{Herm}(d) \mid P > 0, \text{trace}P = 1\}$$

$$T_P \mathcal{S} \cong \mathcal{A}_0 \quad \mathcal{A}_0 = \{X \in \text{Herm}(d) \mid \text{trace}X = 0\}$$

We denote by  $\widetilde{X}$  the corresponding vector field of  $X$ .

For  $P \in \mathcal{S}$ ,  $X \in \mathcal{A}_0$ , define  $\omega_P(\widetilde{X})$  ( $\in \text{Herm}(d)$ ) and  $\xi$  by

$$X = \frac{1}{2}(P\omega_P(\widetilde{X}) + \omega_P(\widetilde{X})P), \quad \xi = -I_d$$

Then  $\{\omega, \xi\}$  is an equiaffine distribution.

(  $\omega_P(\widetilde{X})$  is the symmetric logarithmic derivative of  $X$  ! )

The induced quantities are given by

$$h_P(\widetilde{X}, \widetilde{Y}) = \frac{1}{2}\text{trace} \left( P(\omega_P(\widetilde{X})\omega_P(\widetilde{Y}) + \omega_P(\widetilde{Y})\omega_P(\widetilde{X})) \right),$$

$$\nabla_{\widetilde{X}} \widetilde{Y} = h_P(\widetilde{X}, \widetilde{Y})P - \frac{1}{2}(X\omega_P(\widetilde{Y}) + \omega_P(\widetilde{Y})X).$$

$\{\omega, \xi\}$  : nondegenerate, equiaffine

$v : M \rightarrow R_{n+1}$  is the **conormal map** of  $\{\omega, \xi\}$

$$\begin{array}{c} \iff \\ \stackrel{\text{def}}{=} \end{array} \quad \begin{aligned} \langle v(p), \xi_p \rangle &= 1, \\ \langle v(p), \omega(X_p) \rangle &= 0 \end{aligned}$$

We define a function on  $\Gamma(TM) \times M$  by

$$\rho(X, q) = \langle v(q), \omega(X) \rangle.$$

$\rho$  is called the **geometric pre-divergence** on  $M$ .

	Information geometry	Differential geometry
$\theta$	natural parameters	$\nabla$ -affine coordinates
	exponential arc	$\nabla$ -geodesic
$\eta$	mixture parameters expectation parameters	$\nabla^*$ -affine coordinates
$g$ or $h$	Fisher metric	Riemannian metric affine fundamental form
$T$ or $C$	skewness	cubic form
$\psi$	cumulant generating function free energy	affine hypersurface
$\phi$	entropy	dual map
	Legendre transformation	dual transformation
$D, \rho, \dots$	Kullback-Leibler divergence relative entropy	geometric divergence affine support function
$\rho, \omega, \dots$	prior distribution	volume form
$T, \dots$	bias-correction	Tchebychev vector
	:	:

## Final remarks

パラメータ空間に多様体構造を入れる話を考えた（情報幾何学）

- Geometry for dually flat spaces
  - AdaBoost, U-Boost, ...
  - Modern information theory (LDPC codes, etc. )
  - Linear programming problems
- Bayesian statistics
  - prior distribution  $\iff$  volume form
- Infinite dimensional case (Orlicz space geometry)
  - affine immersion into an functional space
- Quantum version of information geometry
  - statistical manifold admitting torsion
- 標本空間に多様体構造を入れる話は，情報幾何学とよばれていない
  - kernel method